# One-way augmented design

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## 2023 - 11 - 18

One-way ANOVA & pairwise comparison post hoc tests in a non-resolvable augmented design.

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	<pre># (install &amp;) load packages pacman::p_load(     car,     conflicted,     desplot,     emmeans,     ggtext,     lme4,     lmerTest,     multcomp,     multcompView,</pre>	

```
tidyverse)
# handle function conflicts
conflicts_prefer(dplyr::filter)
conflicts_prefer(dplyr::select)
conflicts_prefer(lmerTest::lmer)
```

## 1 Data

This example is taken from Chapter "3.7 Analysis of a non-resolvable augmented design" of the course material "Mixed models for metric data (3402-451)" by Prof. Dr. Hans-Peter Piepho. It considers data published in Petersen (1994) from a yield trial laid out as an augmented design. The genotypes (gen) include 3 standards (st, ci, wa) and 30 new cultivars of interest. The trial was laid out in 6 blocks (block). The 3 standards are tested in each block, while each entry is tested in only one of the blocks. Therefore, the blocks are "incomplete blocks".

#### 1.1 Import

```
# data is available online:
path <- "https://raw.githubusercontent.com/SchmidtPaul/dsfair_quarto/master/data/Petersen1
dat <- read_csv(path) # use path from above</pre>
dat
# A tibble: 48 x 5
          yield block
   gen
                         row
                                col
   <chr> <dbl> <chr> <dbl> <dbl> <dbl>
           2972 I
                            1
 1 \, \text{st}
                                   1
 2 14
           2405 I
                            2
                                   1
                            3
 3 26
           2855 I
                                   1
           2592 I
                            4
                                   1
 4 ci
           2572 I
 5 17
                            5
                                   1
           2608 I
                            6
                                   1
 6 wa
 7 22
           2705 I
                            7
                                   1
 8 13
           2391 I
                            8
                                   1
                            1
                                   2
 9 st
           3122 II
                            2
                                   2
10 ci
           3023 II
# i 38 more rows
```

#### 1.2 Format

Before anything, the columns gen and block should be encoded as factors, since R by default encoded them as character.

```
dat <- dat %>%
  mutate(across(c(gen, block), ~ as.factor(.x)))
```

#### 1.3 Explore

We make use of dlookr::describe() to conveniently obtain descriptive summary tables. Here, we get can summarize per block and per cultivar.

```
dat %>%
 group_by(gen) %>%
 dlookr::describe(yield) %>%
 select(2:sd) %>%
 arrange(desc(n), desc(mean))
```

```
# A tibble: 33 x 5
```

	gen	n	na	mean	sd
	<fct></fct>	<int></int>	< int >	<dbl></dbl>	<dbl></dbl>
1	st	6	0	2759.	832.
2	ci	6	0	2726.	711.
3	wa	6	0	2678.	615.
4	19	1	0	3643	NA
5	11	1	0	3380	NA
6	07	1	0	3265	NA
7	03	1	0	3055	NA
8	04	1	0	3018	NA
9	01	1	0	3013	NA
10	30	1	0	2955	NA
# :	i 23 ma	ore rou	IS		

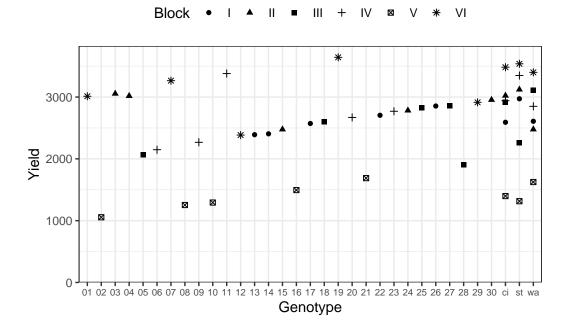
```
dat %>%
group_by(block) %>%
dlookr::describe(yield) %>%
select(2:sd) %>%
arrange(desc(mean))
```

```
# A tibble: 6 x 5
 block
            n
                 na mean
                             sd
  <fct> <int> <int> <dbl> <dbl>
            8
                  0 3205. 417.
1 VI
                  0 2864.
2 II
            8
                           258.
3 IV
            8
                  0 2797. 445.
4 I
            8
                  0 2638. 202.
                  0 2567. 440.
            8
5 III
6 V
            8
                  0 1390.
                           207.
```

Additionally, we can decide to plot our data. Note that we here define custom colors for the genotypes, where all unreplicated entries get a shade of green and all replicated checks get a shade of red.

```
greens30 <- colorRampPalette(c("#bce2cc", "#00923f"))(30)</pre>
oranges3 <- colorRampPalette(c("#e4572e", "#ad0000"))(3)</pre>
gen_cols <- set_names(c(greens30, oranges3), nm = levels(dat$gen))</pre>
ggplot(data = dat) +
  aes(
    y = yield,
    x = gen,
    shape = block
  ) +
  geom point() +
    scale_x_discrete(
    name = "Genotype"
  ) +
  scale_y_continuous(
    name = "Yield",
    limits = c(0, NA),
    expand = expansion(mult = c(0, 0.05))
  ) +
  scale_color_manual(
    guide = "none",
    values = gen_cols
  ) +
  scale_shape_discrete(
    name = "Block"
  ) +
  guides(shape = guide_legend(nrow = 1)) +
```

```
theme_bw() +
theme(
  legend.position = "top",
  axis.text.x = element_text(size = 7)
)
```



Finally, since this is an experiment that was laid with a certain experimental design (= a non-resolvable augmented design) - it makes sense to also get a field plan. This can be done via desplot() from  $\{desplot\}$ .

```
desplot(
   data = dat,
   flip = TRUE, # row 1 on top, not on bottom
   form = gen ~ col + row, # fill color per cultivar
   col.regions = gen_cols, # custom fill colors
   out1 = block, # line between blocks
   text = gen, # cultivar names per plot
   cex = 1, # cultviar names: font size
   shorten = FALSE, # cultivar names: don't abbreviate
   main = "Field layout", # plot title
   show.key = FALSE # hide legend
```

## **Field layout**

st	st	st
14	ci	18
26	04	27
ci	15	ci
17	30	25
wa	03	28
22	wa	05
13	24	wa
st	st	st
09	02	29
06	21	07
Ci	wa	сі
wa	ci	01
20	10	wa
11	08	12
23	16	19

## 2 Model

Finally, we can decide to fit a linear model with yield as the response variable and gen as fixed effects, since our goal is to compare them to each other. Since the trial was laid out in blocks, we also need block effects in the model, but these can be taken either as a fixed or as random effects. Since our goal is to compare genotypes, we will determine which of the two models we prefer by comparing the average standard error of a difference (s.e.d.) for the comparisons between adjusted genotype means - the lower the s.e.d. the better.

```
[1] 462.0431
```

As a result, we find that the model with fixed block effects has the slightly smaller s.e.d. and is therefore more precise in terms of comparing genotypes.

```
▲ Model assumptions met? (click to show)
```

It would be at this moment (i.e. after fitting the model and before running the ANOVA), that you should check whether the model assumptions are met. Find out more in the summary article "Model Diagnostics"

## 3 ANOVA

Based on our model, we can then conduct an ANOVA:

```
ANOVA <- car::Anova(mod_fb, type = "III")
ANOVA
```

Anova Table (Type III tests)

```
Response: yield
             Sum Sq Df F value
                                  Pr(>F)
(Intercept) 3073607
                    1
                        33.738 0.0001710 ***
            4095905 32
                         1.405 0.2930113
gen
                       15.298 0.0002082 ***
block
            6968486 5
             911027 10
Residuals
____
                0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Signif. codes:
```

Accordingly, the ANOVA's F-test found the cultivar effects to be statistically significant (p = 0.293). Additionally, the block effects are also statistically significant (p <  $.001^{***}$ ), but this is only of secondary concern for us.

#### 4 Mean comparison

80

2528 341 10

1060

Besides an ANOVA, one may also want to compare adjusted yield means between cultivars via post hoc tests (t-test, Tukey test etc.).

```
mean_comp <- mod_fb %>%
  emmeans(specs = ~ gen) %>% # adj. mean per genotype
  cld(adjust = "Tukey", Letters = letters) # compact letter display (CLD)
mean_comp
 gen emmean SE df lower.CL upper.CL .group
 12
       1632 341 10
                        164
                                3100 a
 06
       1823 341 10
                        355
                                3291 a
 28
       1862 341 10
                        394
                                3330 a
 09
       1943 341 10
                        475
                                3411 a
 05
       2024 341 10
                        556
                                3492 a
 29
       2162 341 10
                        694
                                3630 a
       2260 341 10
 01
                        792
                                3728
                                      а
 15
       2324 341 10
                                3792 a
                        856
 02
       2330 341 10
                        862
                                3798 a
 20
       2345 341 10
                        877
                                3813 a
 13
       2388 341 10
                        920
                                3856
                                      а
 14
       2402 341 10
                        934
                                3870 a
 23
       2445 341 10
                        977
                                3913
                                       а
 07
       2512 341 10
                       1044
                                3980
                                       а
```

а

3996

18	2562	341	10	1094	4030	a
10	2568	341	10	1100	4036	a
17	2569	341	10	1101	4037	a
24	2630	341	10	1162	4098	a
wa	2678	123	10	2148	3208	a
22	2702	341	10	1234	4170	a
ci	2726	123	10	2195	3256	a
st	2759	123	10	2229	3289	a
16	2770	341	10	1302	4238	a
25	2784	341	10	1316	4252	a
30	2802	341	10	1334	4270	a
27	2816	341	10	1348	4284	a
26	2852	341	10	1384	4320	a
04	2865	341	10	1397	4333	a
19	2890	341	10	1422	4358	a
03	2902	341	10	1434	4370	a
21	2963	341	10	1495	4431	a
11	3055	341	10	1587	4523	a

```
Results are averaged over the levels of: block
Confidence level used: 0.95
Conf-level adjustment: sidak method for 33 estimates
P value adjustment: tukey method for comparing a family of 33 estimates
significance level used: alpha = 0.05
NOTE: If two or more means share the same grouping symbol,
        then we cannot show them to be different.
        But we also did not show them to be the same.
```

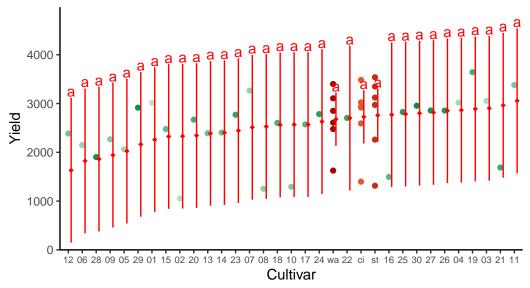
It can be seen that while some genotypes have a higher yield than others, no differences are found to be statistically significant here. Accordingly, notice that *e.g.* for gen 11, which is the genotype with the highest adjusted yield mean (=3055), its lower confidence limit (=1587) includes gen 12, which is the genotype with the lowest adjusted yield mean (=1632).

Note that if you would like to see the underlying individual contrasts/differences between adjusted means, simply add details = TRUE to the cld() statement. Furthermore, check out the Summary Article "Compact Letter Display".

Finally, we can create a plot that displays both the raw data and the results, *i.e.* the comparisons of the adjusted means that are based on the linear model.

```
# reorder genotype factor levels according to adjusted mean
my_caption <- "Dots represent raw data. Red diamonds and error bars represent adjusted mea</pre>
```

```
ggplot() +
  # green/red dots representing the raw data
 geom_point(
    data = dat,
    aes(y = yield, x = gen, color = gen)
  ) +
  # red diamonds representing the adjusted means
  geom_point(
   data = mean_comp,
   aes(y = emmean, x = gen),
   shape = 18,
   color = "red",
    position = position_nudge(x = 0.2)
  ) +
  # red error bars representing the confidence limits of the adjusted means
  geom_errorbar(
   data = mean_comp,
    aes(ymin = lower.CL, ymax = upper.CL, x = gen),
    color = "red",
   width = 0.1,
   position = position_nudge(x = 0.2)
  ) +
  # red letters
  geom_text(
   data = mean_comp,
   aes(y = upper.CL, x = gen, label = str_trim(.group)),
   color = "red",
   vjust = -0.2,
   position = position_nudge(x = 0.2)
  ) +
  scale_color_manual(
   guide = "none",
   values = gen_cols
  ) +
  scale_x_discrete(
   name = "Cultivar",
   limits = as.character(mean_comp$gen)
  ) +
  scale_y_continuous(
   name = "Yield",
   limits = c(0, NA),
```



Dots represent raw data. Red diamonds and error bars represent adjusted means with 95% confidence limits per cultivar. Means followed by a common letter are not significantly different according to the Tukey-test.

### 5 Bonus

Here are some other things you would maybe want to look at for the analysis of this dataset.

#### 5.1 Variance components

To extract variance components from our models, we unfortunately need different functions per model since only of of them is a mixed model and we used different functions to fit them.

```
# Residual Variance
summary(mod_fb)$sigma<sup>2</sup>
```

[1] 91102.66 # Both Variance Components as\_tibble(VarCorr(mod\_rb)) # A tibble: 2 x 5 grp var1 var2 vcov sdcor <chr> <dbl> <dbl> <chr> <chr> 1 block (Intercept) <NA> 434198. 659. 2 Residual <NA> 302. <NA> 91103.

Petersen, Roger G. 1994. Agricultural Field Experiments. CRC Press. https://doi.org/10. 1201/9781482277371.